



# Combine Tutorial

Terascale Statistics School  
Kyle Cormier, Aliya Namigova, Nick Wardle

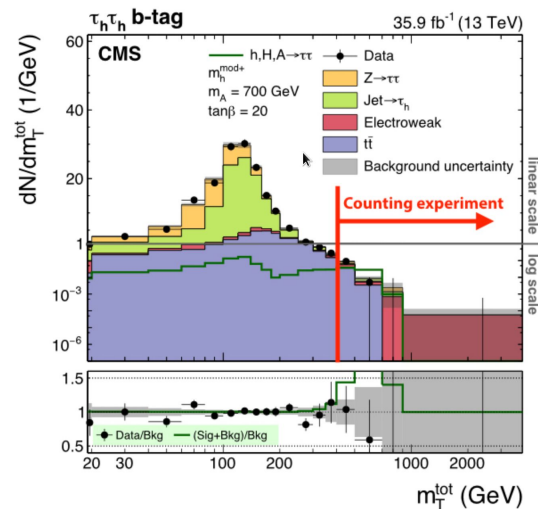
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# Introduction

# Overview

Search for a heavy neutral higgs,  
Which decays to  $\tau\tau$

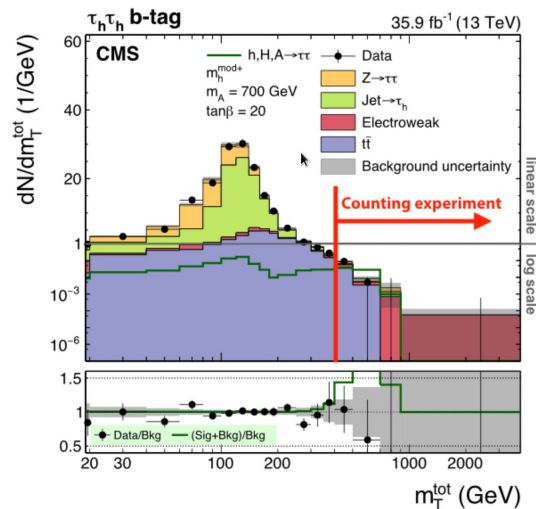
1. Simple Counting Experiment
2. Shape Based Analysis
3. Adding Control Regions
4. Physics models – beyond a single  
Signal strength parameter



# Overview

Search for a heavy neutral higgs,  
Which decays to  $\tau\tau$

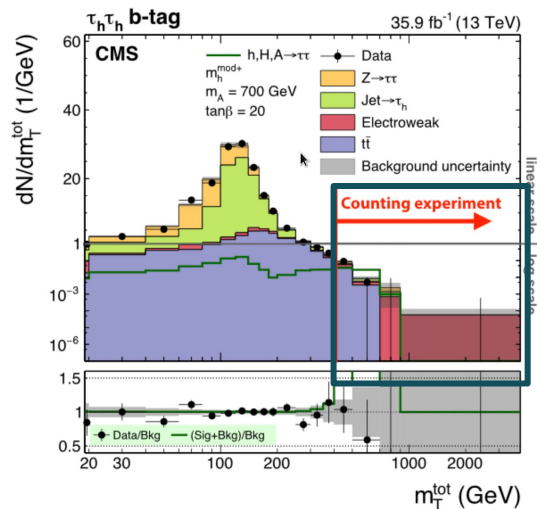
- Limit Setting
- Significance Testing
- Asymptotic Calculations
- Toy-based Calculations
- Parameter Extraction
- Fit Debugging
- Fit Plotting
- Nuisance Parameter Checks
- Multi-Dimensional Likelihood Scans



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# Part 1 – Counting Experiment

# Counting Experiment



Count number of events in high  $m_T$  region

imax 1 number of bins

jmax 4 number of processes minus 1

kmax \* number of nuisance parameters

----- observation -----

bin signal\_region

observation 10.0

backgrounds

signal

-----

bin signal\_region signal\_region signal\_region signal\_region signal\_region

process ttbar diboson Ztautau jetFakes bbHtautau

process 1 2 3 4 0

rate 4.43803 3.18309 3.7804 1.63396 0.711064

-----

CMS\_eff\_b lnN 1.02 1.02 1.02 - 1.02

CMS\_eff\_t lnN 1.12 1.12 1.12 - 1.12

CMS\_eff\_t\_highpt lnN 1.1 1.1 1.1 - 1.1

acceptance\_Ztautau lnN - - 1.08 - -

acceptance\_bbH lnN - - - - 1.05

acceptance\_ttbar lnN 1.005 - - - -

norm\_jetFakes lnN - - - 1.2 -

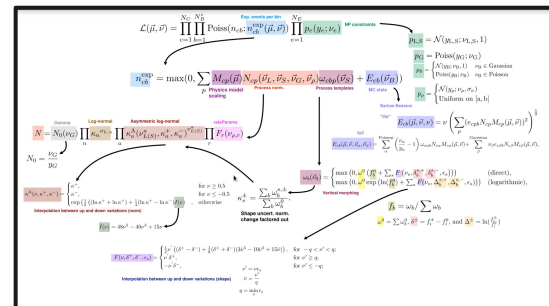
xsec\_diboson lnN - 1.05 - - -

Systematic uncertainties

## Neyman-Pearson doesn't usually help

We usually don't have explicit formulae for the pdfs  $f(x|s)$ ,  $f(x|b)$ , so for a given  $x$  we can't evaluate the likelihood ratio

$$t(x) = \frac{f(x|s)}{f(x|b)}$$



## Use a modified version of the likelihood ratio test statistic

### Test statistic based on likelihood ratio

How can we choose a test's critical region in an 'optimal way', in particular if the data space is multidimensional?

Neyman-Pearson lemma states:

For a test of  $H_0$  of size  $\alpha$ , to get the highest power with respect to the alternative  $H_1$  we need for all  $x$  in the critical region  $W$

"likelihood ratio (LR)"  $\rightarrow \frac{P(x|H_1)}{P(x|H_0)} \geq c_\alpha$

inside  $W$  and  $\leq c_\alpha$  outside, where  $c_\alpha$  is a constant test of the desired size.

Equivalently, optimal scalar test statistic is

N.B. any monotonic function of this is lead

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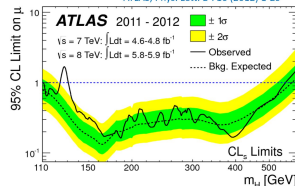
### How to read the green and yellow limit plots

For every value of  $m_{H^\pm}$ , find the upper limit on  $\mu$ .

Also for each  $m_{H^\pm}$ , determine the distribution of upper limits  $\mu_{up}$  one would obtain under the hypothesis of  $\mu = 0$ .

The dashed curve is the median  $\mu_{up}$ , and the green (yellow) bands give the  $\pm 1\sigma$  ( $2\sigma$ ) regions of this distribution.

ATLAS, Phys. Lett. B 716 (2012) 1-29



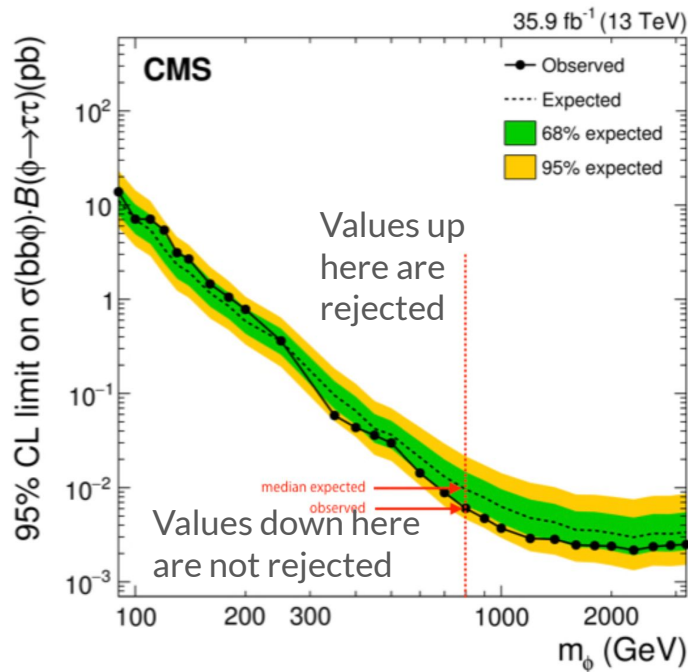
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We use:

$$\tilde{q}_\mu = \begin{cases} -2 \log \left( \frac{\mathcal{L}(\mu)}{\mathcal{L}(\mu=0)} \right) & \hat{\mu} < 0 \\ -2 \log \left( \frac{\mathcal{L}(\mu)}{\mathcal{L}(\hat{\mu})} \right) & 0 < \hat{\mu} < \mu \\ 0 & \mu < \hat{\mu} \end{cases} \quad (\text{And use CL}_s \text{ criterion})$$



# Asymptotic vs Using Toys

## Wilks' Theorem

Wilks' Theorem: if the hypothesized  $\mu_i(\theta)$ ,  $i = 1, \dots, N$ , are true for some choice of the parameters  $\theta = (\theta_1, \dots, \theta_M)$ , then in the large sample limit (and provided regularity conditions are satisfied)

$$t_\mu = -2 \ln \frac{L(\boldsymbol{\mu}(\hat{\boldsymbol{\theta}}))}{L(\hat{\boldsymbol{\mu}})}$$

MLE of  $(\theta_1, \dots, \theta_M)$  follows a chi-square distribution for  $N - M$  degrees of freedom.

MLE of  $(\mu_1, \dots, \mu_N)$

The regularity conditions include: the model in the numerator of the likelihood ratio is “nested” within the one in the denominator, i.e.,  $\boldsymbol{\mu}(\theta)$  is a special case of  $\boldsymbol{\mu} = (\mu_1, \dots, \mu_N)$ .

Proof boils down to having all estimators  $\sim$  Gaussian.

S.S. Wilks, *The large-sample distribution of the likelihood ratio for testing composite hypotheses*, Ann. Math. Statist. **9** (1938) 60-2.

Under some conditions the distribution of the test statistic is known analytically:

→ Use asymptotic approximation

Otherwise:

Generate many sets of pseudodata to get an empirical distribution of the test statistic



## Asymptotic Approximation

```
<<< Combine >>>
<<< v9.1.0 >>>
>>> Random number generator seed is 123456
>>> Method used is AsymptoticLimits

-- AsymptoticLimits ( CLs ) --
Observed Limit: r < 10.8183
Expected 2.5%: r < 7.0537
Expected 16.0%: r < 9.8108
Expected 50.0%: r < 14.5625
Expected 84.0%: r < 22.3988
Expected 97.5%: r < 33.5971
```

## Empirical Distribution

```
-- Hybrid New --
Limit: r < 11.1291 +/- 0.163054 @ 95% CL
```

## Expected

2.5%	Limit: r < 5.46875 +/- 0.15625 @ 95% CL
16.0%	Limit: r < 10.4676 +/- 0.123997 @ 95% CL
50.0%	Limit: r < 14.5396 +/- 0.136762 @ 95% CL
84.0%	Limit: r < 21.7222 +/- 0.271188 @ 95% CL
97.5%	Limit: r < 33.2392 +/- 1.62741 @ 95% CL

# Empirical test statistic distributions

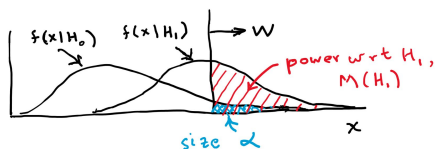
Can directly look at the distributions of the test statistics under the background-only and signal\_background hypothesis

## Definition of a test (2)

But in general there are an infinite number of possible critical regions that give the same size  $\alpha$ .

Use the alternative hypothesis  $H_1$  to motivate where to place the critical region.

Roughly speaking, place the critical region where there is a low probability ( $\alpha$ ) to be found if  $H_0$  is true, but high if  $H_1$  is true:



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Cowan, Cranmer, Gross, Vitells, arXiv:1007.1727, EPIC 71 (2011) 1554

## Monte Carlo test of asymptotic formulae

Consider again  $n \sim \text{Poisson}(\mu s + b)$ ,  $m \sim \text{Poisson}(\tau b)$

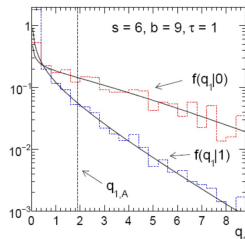
Use  $q_\mu$  to find  $p$ -value of hypothesized  $\mu$  values.

E.g.  $f(q_1|1)$  for  $p$ -value of  $\mu=1$ .

Typically interested in 95% CL, i.e.,  $p$ -value threshold = 0.05, i.e.,  $q_1 = 2.69$  or  $Z_1 = \sqrt{q_1} = 1.64$ .

Median  $[q_1|0]$  gives "exclusion sensitivity".

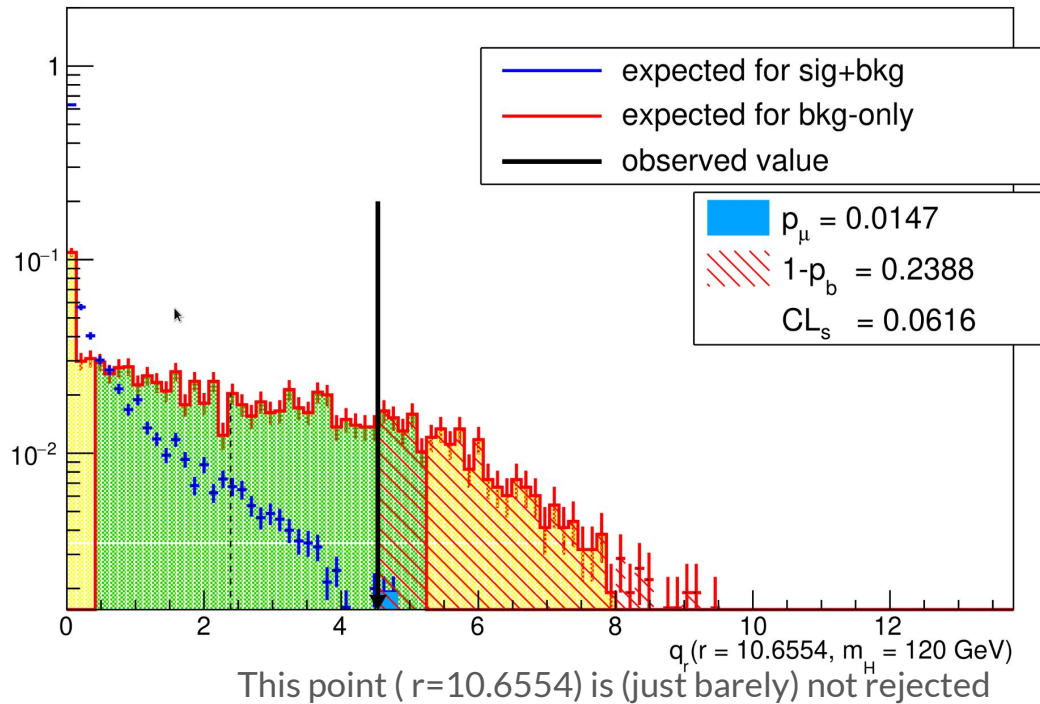
Here asymptotic formulae good for  $s=6$ ,  $b=9$ .



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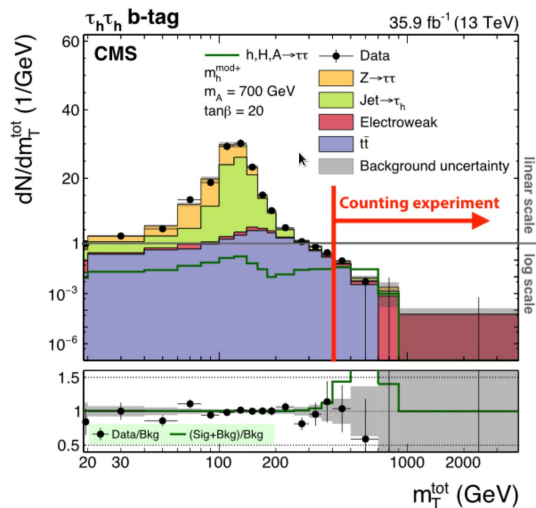
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# Part 2 - Shape Experiment



Use the full histogram!

```
imax 1
jmax 1
kmax *
```

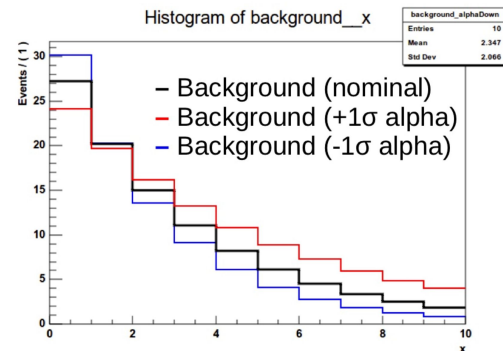
## Input histograms

```
shapes * * simple-shapes-TH1_input.root $PROCESS $PROCESS_$SYSTEMATIC
shapes signal * simple-shapes-TH1_input.root $PROCESS$MASS $PROCESS$MASS_$SYSTEMATIC
```

```
bin bin1
observation 85
```

bin	bin1	bin1
process	signal	background
process	0	1
rate	10	100

lumi	lnN	1.10	1.0
bgnorm	lnN	1.00	1.3
alpha	shape	-	1



```

imax      1 number of bins
jmax      4 number of processes minus 1
kmax      * number of nuisance parameters
-----
shapes bbHtautau * datacard_part2.shapes.root signal_region/$PROCESS$MASS signal_region/$PROCESS$MASS_$SYSTEMATIC
shapes * * datacard_part2.shapes.root signal_region/$PROCESS signal_region/$PROCESS_$SYSTEMATIC
-----
bin          signal_region
observation  3416.0
-----
bin          signal_region  signal_region  signal_region  signal_region  signal_region
process      ttbar          diboson       Ztautau       jetFakes       bbHtautau
process      1              2              3              4              0
rate         683.017        96.5185       742.649       2048.94        0.913183
-----
CMS_eff_b    lnN          1.02          1.02          1.02          -              1.02
CMS_eff_t    lnN          1.12          1.12          1.12          -              1.12
acceptance_Ztautau lnN          -              -              1.08          -              -
acceptance_bbH   lnN          -              -              -              -              1.05
acceptance_ttbar lnN          1.005         -              -              -              -
lumi_13TeV     lnN          1.025         1.025         1.025         -              1.025
norm_jetFakes  lnN          -              -              -              1.2           -
xsec_Ztautau   lnN          -              -              1.04          -              -
xsec_diboson   lnN          -              1.05          -              -              -
xsec_ttbar     lnN          1.06          -              -              -              -
# These ones are new
top_pt_ttbar_shape shape      1              -              -              -              -
CMS_scale_t_1prong0pi0_13TeV shape      1              1              1              -              1
CMS_scale_t_1prong1pi0_13TeV shape      1              1              1              -              1
CMS_scale_t_3prong0pi0_13TeV shape      1              1              1              -              1
CMS_eff_t_highpt shape      1              1              1              -              1

```

## Shape-based analysis improved limits over simple counting experiment

```
-- AsymptoticLimits ( CLs ) --  
Observed Limit: r < 7.9771  
Expected 2.5%: r < 4.7720  
Expected 16.0%: r < 6.8417  
Expected 50.0%: r < 10.5312  
Expected 84.0%: r < 16.9959  
Expected 97.5%: r < 26.5059
```

Shape-based

```
<<< Combine >>>  
<<< v9.1.0 >>>  
>>> Random number generator seed is 123456  
>>> Method used is AsymptoticLimits  
  
-- AsymptoticLimits ( CLs ) --  
Observed Limit: r < 10.8183  
Expected 2.5%: r < 7.0537  
Expected 16.0%: r < 9.8108  
Expected 50.0%: r < 14.5625  
Expected 84.0%: r < 22.3988  
Expected 97.5%: r < 33.5971
```

Counting

# Fit diagnostics

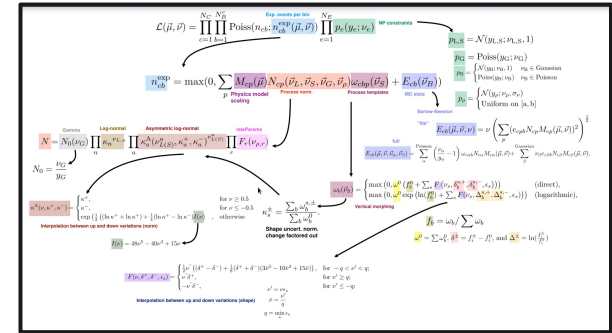
## Neyman-Pearson doesn't usually help

We usually don't have explicit formulae for the pdfs  $f(\mathbf{x}|s)$ ,  $f(\mathbf{x}|b)$ , so for a given  $\mathbf{x}$  we can't evaluate the likelihood ratio

$$t(\mathbf{x}) = \frac{f(\mathbf{x}|s)}{f(\mathbf{x}|b)}$$

In this case we do have explicit formulas for the pdf, we constructed it with combine!

BUT we want to make sure the model we've constructed is sensible and the fits are running well! → **Inspect various aspects of your model!**



## Finally

Two lectures only enough for a brief introduction to:

Parameter estimation

Hypothesis tests (→ path to Machine Learning)

Limits (confidence intervals/regions)

Systematics (nuisance parameters)

No time for many other interesting topics:

Experimental sensitivity

Bayesian parameter estimation

Final thought: once the basic formalism is fixed, most of the work focuses on writing down the likelihood, e.g.,  $P(\mathbf{x}|\theta)$ , and including in it enough parameters to adequately describe the data (true for both Bayesian and frequentist approaches) so often best to invest most of your time with it.

# Fit Parameter Values

RooFitResult: minimized FCN value: -2.55338e-05, estimated distance to minimum: 7.54243e-06  
covariance matrix quality: Full, accurate covariance matrix  
Status : MINIMIZE=0 HESSE=0

Floating Parameter	FinalValue +/-	Error
CMS_eff_b	-4.5380e-02 +/-	9.93e-01
CMS_eff_t	-2.6311e-01 +/-	7.33e-01
CMS_eff_t_highpt	-4.7146e-01 +/-	9.62e-01
CMS_scale_t_1prong0pi0_13TeV	-1.5989e-01 +/-	5.93e-01
CMS_scale_t_1prong1pi0_13TeV	-1.6426e-01 +/-	4.94e-01
CMS_scale_t_3prong0pi0_13TeV	-3.0698e-01 +/-	6.06e-01
acceptance_Ztautau	-3.1262e-01 +/-	8.62e-01
acceptance_bbH	-2.8676e-05 +/-	1.00e+00
acceptance_ttbar	4.9981e-03 +/-	1.00e+00
lumi_13TeV	-5.6366e-02 +/-	9.89e-01
norm_jetFakes	-9.3327e-02 +/-	2.56e-01
r	-2.7220e+00 +/-	2.59e+00
top_pt_ttbar_shape	1.7586e-01 +/-	7.00e-01
xsec_Ztautau	-1.6007e-01 +/-	9.66e-01
xsec_diboson	3.9758e-02 +/-	1.00e+00
xsec_ttbar	5.7794e-02 +/-	9.46e-01



name	b-only fit	s+b fit	rho
CMS_eff_b	-0.04, 0.99	-0.05, 0.99	+0.01
CMS_eff_t	* -0.24, 0.73*	* -0.26, 0.73*	+0.06
CMS_eff_t_highpt	* -0.56, 0.94*	* -0.47, 0.96*	+0.02
CMS_scale_t_1prong0pi0_13TeV	* -0.17, 0.58*	* -0.16, 0.59*	-0.04
CMS_scale_t_1prong1pi0_13TeV	! -0.12, 0.45!	! -0.16, 0.49!	+0.20
CMS_scale_t_3prong0pi0_13TeV	* -0.31, 0.61*	* -0.31, 0.61*	+0.02
acceptance_Ztautau	* -0.31, 0.86*	* -0.31, 0.86*	-0.05
acceptance_bbH	+0.00, 1.00	-0.00, 1.00	+0.05
acceptance_ttbar	+0.01, 1.00	+0.00, 1.00	+0.00
lumi_13TeV	-0.05, 0.99	-0.06, 0.99	+0.01
norm_jetFakes	! -0.09, 0.26!	! -0.09, 0.26!	-0.05
top_pt_ttbar_shape	* +0.24, 0.69*	* +0.18, 0.70*	+0.22
xsec_Ztautau	-0.16, 0.97	-0.16, 0.97	-0.02
xsec_diboson	+0.03, 1.00	+0.04, 1.00	-0.02
xsec_ttbar	+0.08, 0.95	+0.06, 0.95	+0.02



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# Part 3

# Rate Parameters

Can add rate parameters which scale certain processes

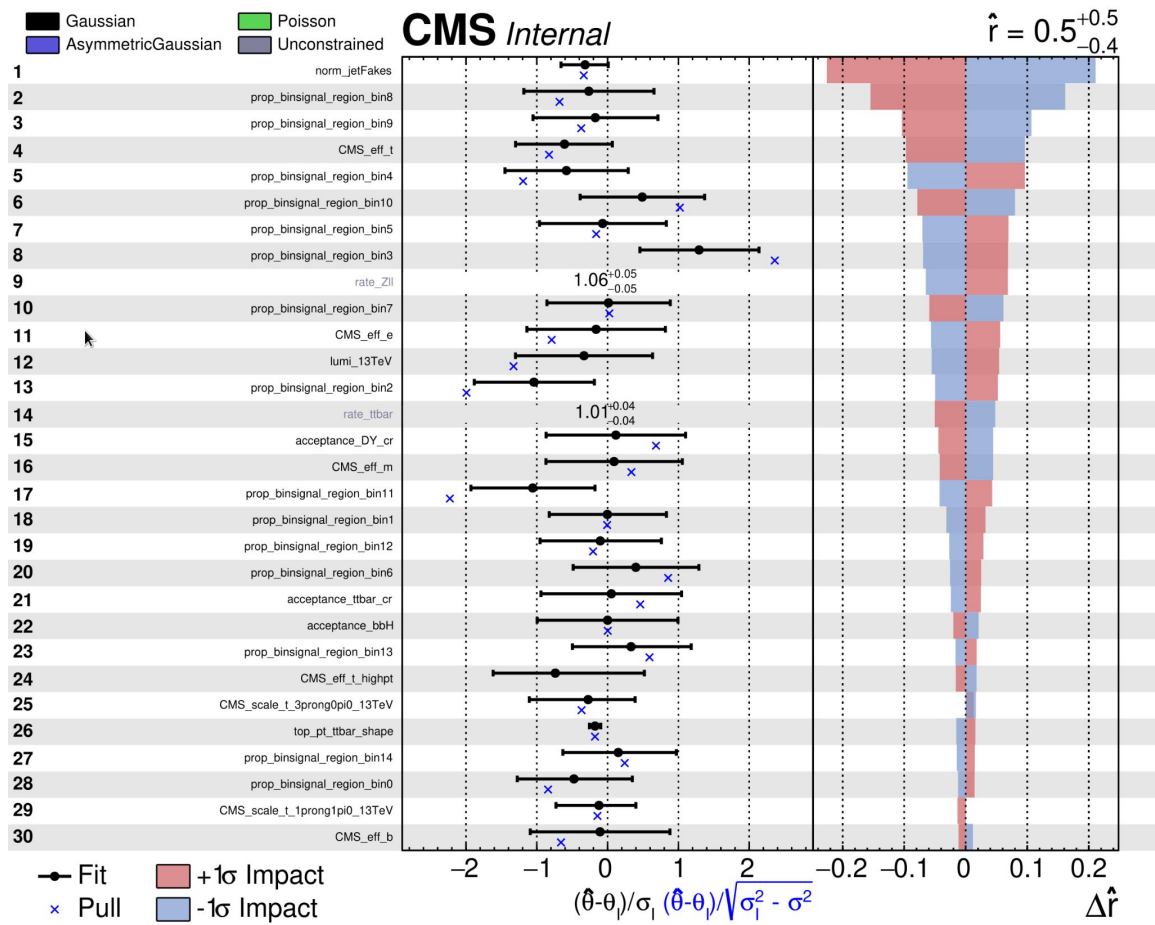
```
[name] rateParam [channel] [process] [init] [min,max]
```



To allow the rates of the ttbar and Z->ll process in the control regions to influence those in the signal region,  
Connect them to each other via a rate parameter

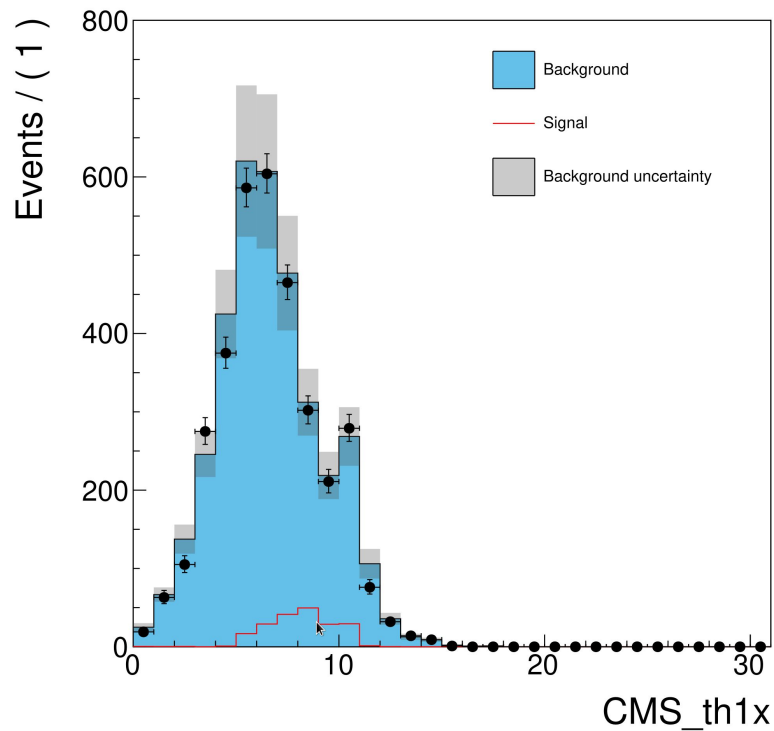
```
rate_ttbar rateParam * ttbar 1  
rate_Zll rateParam * Ztautau 1  
rate_Zll rateParam * Zmumu 1
```

# Impacts

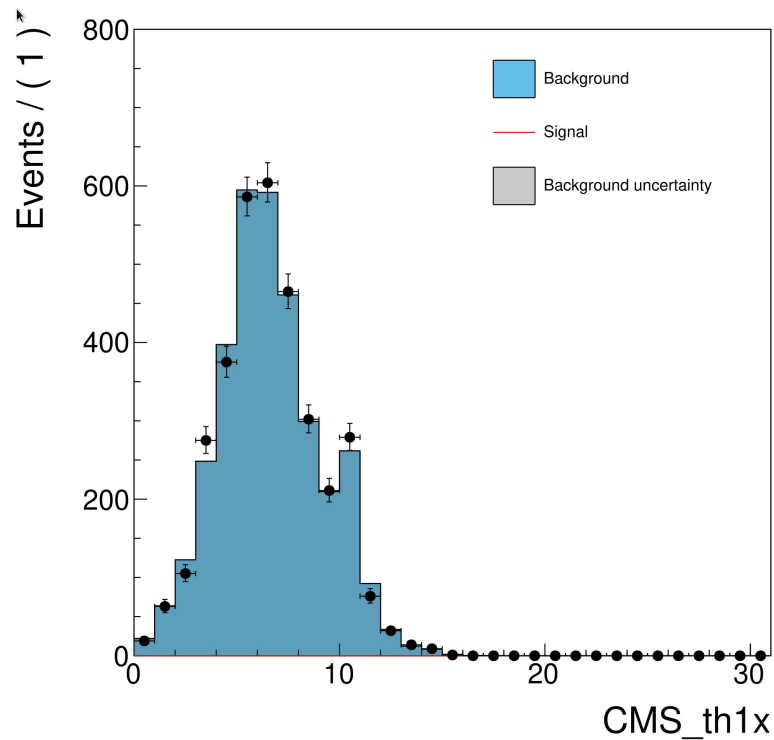


# Visualizing fits

Prefit



background -only fit



# Significance

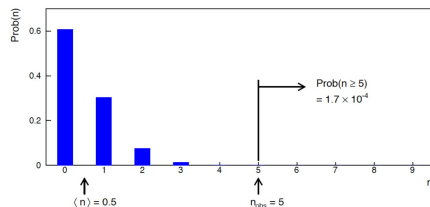
## Poisson counting experiment: discovery $p$ -value

Suppose  $b = 0.5$  (known), and we observe  $n_{\text{obs}} = 5$ .

Should we claim evidence for a new discovery?

Give  $p$ -value for hypothesis  $s = 0$ , suppose relevant alt. is  $s > 0$ .

$$\begin{aligned} p\text{-value} &= P(n \geq 5; b = 0.5, s = 0) \\ &= 1.7 \times 10^{-4} \neq P(s = 0)! \end{aligned}$$



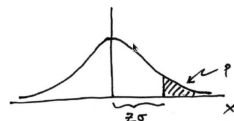
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## Significance from $p$ -value

Often define significance  $Z$  as the number of standard deviations that a Gaussian variable would fluctuate in one direction to give the same  $p$ -value.



$$p = \int_Z^{\infty} \frac{1}{\sqrt{2\pi}} e^{-x^2/2} dx = 1 - \Phi(Z)$$

$$Z = \Phi^{-1}(1 - p)$$

in ROOT:

```
p = 1 - TMath::Freq(Z)
Z = TMath::NormQuantile(1-p)
```

in python (scipy.stats):

```
p = 1 - norm.cdf(Z) = norm.sf(Z)
Z = norm.ppf(1-p)
```

Result  $Z$  is a “number of sigmas”. Note this does not mean that the original data was Gaussian distributed.

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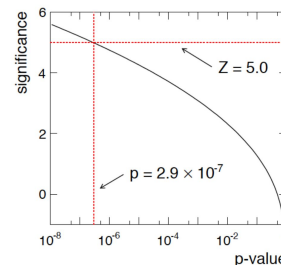
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## Poisson counting experiment: discovery significance

Equivalent significance for  $p = 1.7 \times 10^{-4}$ :  $Z = \Phi^{-1}(1 - p) = 3.6$

Often claim discovery if  $Z > 5$  ( $p < 2.9 \times 10^{-7}$ , i.e., a “5-sigma effect”)



In fact this tradition should be revisited:  $p$ -value intended to quantify probability of a signal-like fluctuation assuming background only; not intended to cover, e.g., hidden systematics, plausibility signal model, compatibility of data with signal, “look-elsewhere effect” (~multiple testing), etc.

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We calculate the  $p$ -value of the modified Likelihood Ratio test-statistic  $q_0$  and quote a significance

$$q_0 = \begin{cases} 0 & \hat{\mu} < 0 \\ -2 \log \left( \frac{\mathcal{L}(\mu_{\text{NP}}=0)}{\mathcal{L}(\hat{\mu}_{\text{NP}})} \right) & \hat{\mu} \geq 0 \end{cases}$$

# Significance

Simple Asymptotic Calculation – assume known distribution of test-statistic

```
<<< Combine >>>
<<< v9.1.0 >>>
>>> Random number generator seed is 123456
>>> Method used is Significance

-- Significance --
Significance: 1.11273
```

Can also instruct combine to give the p-value directly with the `--pvalue` flag

```
-- Significance --
p-value of background: 0.132912
Done in 0.00 min (cpu), 0.00 min (real)
```

Asymptotic method using the large sample limit (+ some other conditions, like Wilk's theorem)

Cowan, Cranmer, Gross, Vitells, arXiv:1007.1727, EPJC 71 (2011) 1554

## Distribution of $q_0$ in large-sample limit

Assuming approximations valid in the large sample (asymptotic) limit, we can write down the full distribution of  $q_0$  as

$$f(q_0|\mu') = \left(1 - \Phi\left(\frac{\mu'}{\sigma}\right)\right) \delta(q_0) + \frac{1}{2} \frac{1}{\sqrt{2\pi}} \frac{1}{\sqrt{q_0}} \exp\left[-\frac{1}{2}\left(\sqrt{q_0} - \frac{\mu'}{\sigma}\right)^2\right]$$

The special case  $\mu' = 0$  is a “half chi-square” distribution:

$$f(q_0|0) = \frac{1}{2} \delta(q_0) + \frac{1}{2} \frac{1}{\sqrt{2\pi}} \frac{1}{\sqrt{q_0}} e^{-q_0/2}$$

In large sample limit,  $f(q_0|0)$  independent of nuisance parameters;  $f(q_0|\mu')$  depends on nuisance parameters through  $\sigma$ .

# Significance

Can also check expected significance for various signal strengths, e.g. with `-t -1 --expectSignal 1.5`

```
-- Significance --  
Significance: 3.52007
```

Can also use fit model after a first fit to the data to get model parameters with `--toysFrequentist`

```
-- Significance --  
Significance: 3.13954
```



# Signal Strength measurement

## Maximum Likelihood Estimators (MLEs)

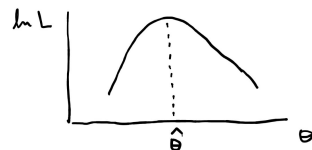
We *define* the maximum likelihood estimators or MLEs to be the parameter values for which the likelihood is maximum.

Maximizing  $L$   
equivalent to  
maximizing  $\log L$

$$\hat{\theta} = \operatorname{argmax}_{\theta} L(\theta)$$

Could have multiple maxima (take highest).

MLEs not guaranteed to have any 'optimal' properties, (but in practice they're very good).



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## Example of variance by graphical method

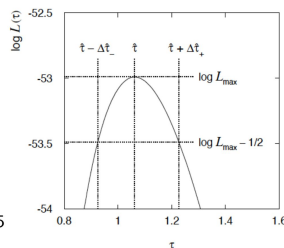
ML example with exponential:

$$\hat{\tau} = 1.062$$

$$\Delta \hat{\tau}_{-} = 0.137$$

$$\Delta \hat{\tau}_{+} = 0.165$$

$$\hat{\sigma}_{\hat{\tau}} \approx \Delta \hat{\tau}_{-} \approx \Delta \hat{\tau}_{+} \approx 0.15$$

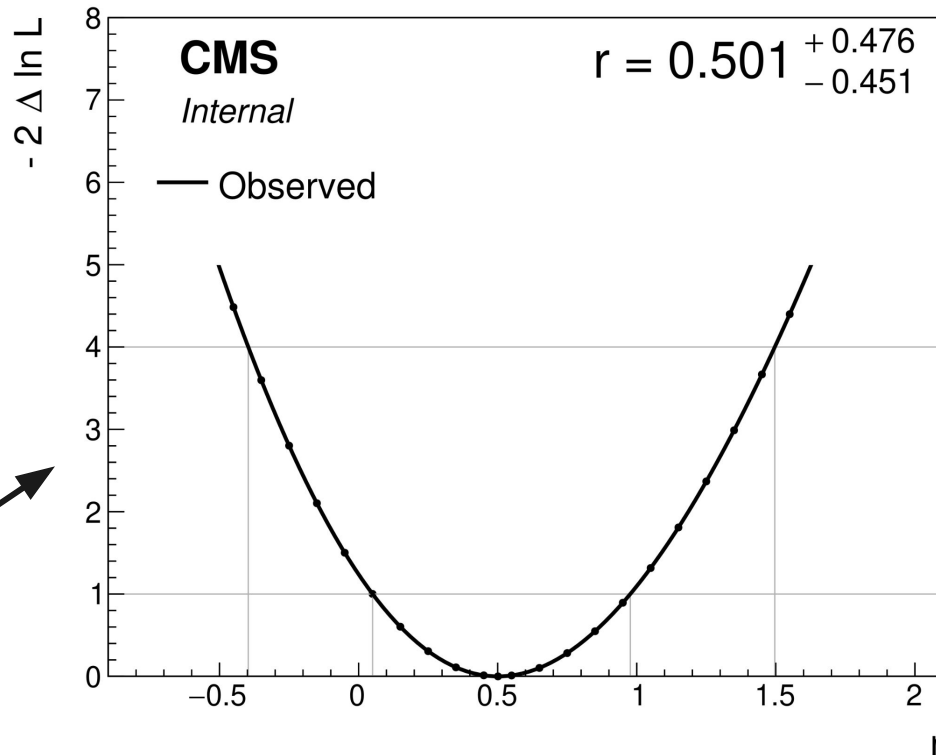


Not quite parabolic  $\ln L$  since finite sample size ( $n = 50$ ).

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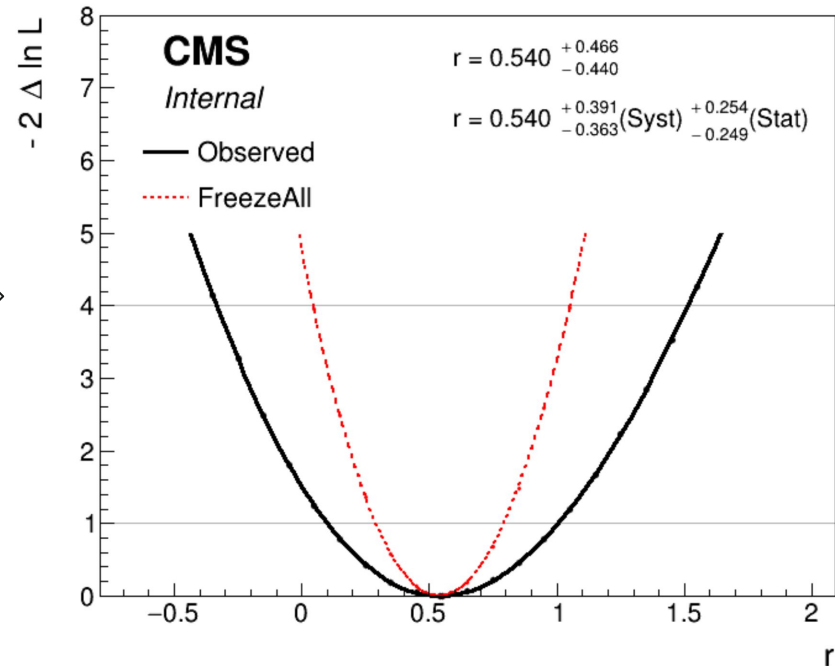
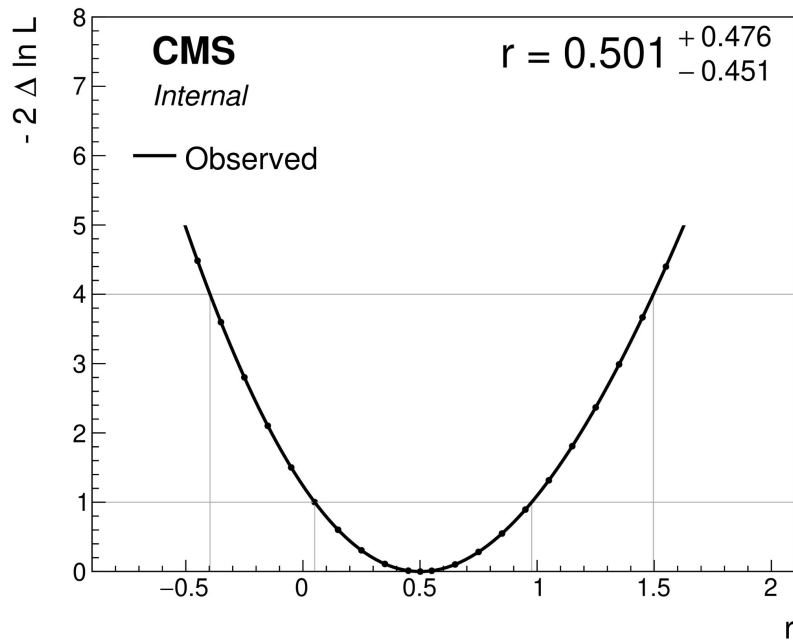
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With a  
sign flip

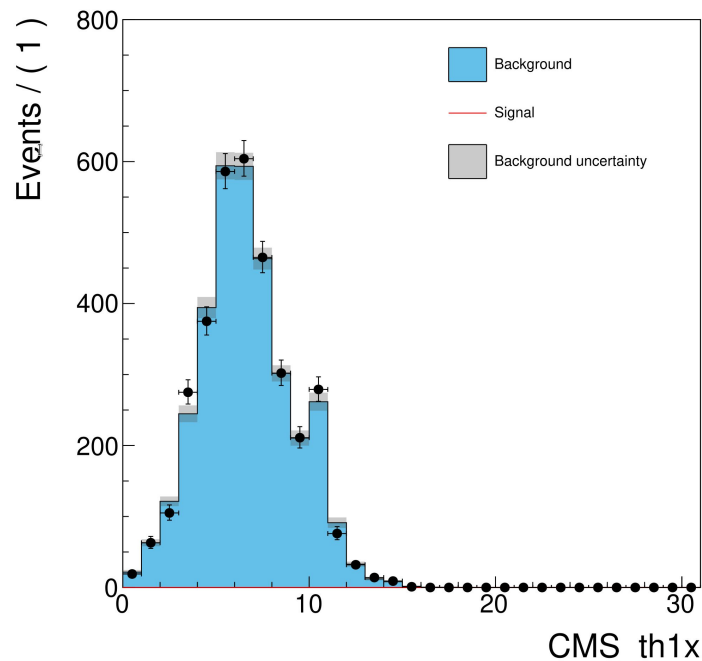
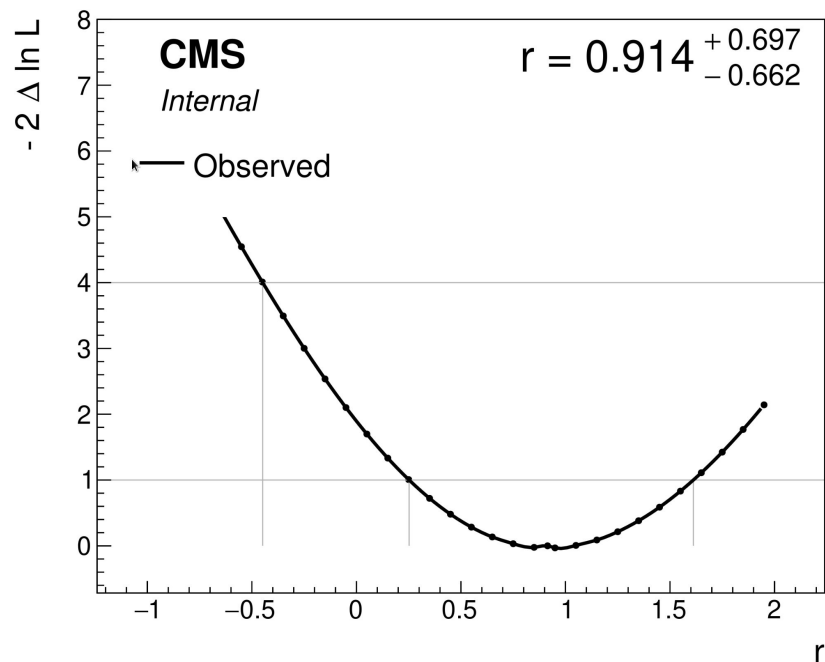


# Uncertainty Breakdowns



# Channel Masks

Can also mask particular channels to investigate the fit, e.g. masking the control regions:



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# Part 4

# Two Parameters of Interest

```
from HiggsAnalysis.CombinedLimit.PhysicsModel import PhysicsModel
```

```
class DASModel(PhysicsModel):
```

```
    def doParametersOfInterest(self):
```

```
        """Create POI and other parameters, and define the POI set."""
```

```
        self.modelBuilder.doVar("r[0,0,10]")
```

```
        self.modelBuilder.doSet("POI", "r".join(["r"]))
```

```
    def getYieldScale(self, bin, process):
```

```
        "Return the name of a RooAbsReal to scale this yield by or the two special values 1 and 0
```

```
        if self.DC.isSignal[process]:
```

```
            print("Scaling %s/%s by r" % (bin, process))
```

```
            return "r"
```

```
        return 1
```

```
dasModel = DASModel()
```

## 2-Dimensional NLL map

