

# Advanced Statistics 2

## Setting upper limits

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# The general scenario

An experiment is searching for

- Charged lepton flavour violation
- A particular supersymmetry decay channel
- Neutrinoless double beta decay
- *insert your topic here*

It sees nothing. Or perhaps a signal so small it could come from background processes.

What can you say?

## A basic case

to set the scene

Suppose zero background. Plan to measure number of events  $N$ . Convert to physics  $R$  (cross-section, decay-rate, wimp density...) by  $R = AN$

Run experiment - get  $N = 0$

Poisson formula is  $P(r; \lambda) = e^{-\lambda} \frac{\lambda^r}{r!}$

What can you say about  $\lambda$ ? Large  $\lambda$  implausible as  $P(0, \lambda)$  very small.

**Frequentist** says:

I note that  $P(0; 2.3) = 0.10$ .

I quote  $\lambda_{UL} = 2.3$  as the upper limit at 90% confidence

**Bayesian** says:

I take prior flat in  $\lambda$ :  $P(\lambda) = \text{const}$ .

The posterior is

$P'(\lambda) = P(N; \lambda) \times \text{const} / \text{const}'$ , so for 0 events,  $P'(\lambda) = e^{-\lambda}$

Most likely value is 0, but I note that  $\int_0^{2.3} e^{-\lambda} d\lambda = 0.9$ .

I quote  $\lambda_{UL} = 2.3$  as the upper limit at 90% confidence

# More details

## on the Frequentist

They said:


"I note that  $P(0; 2.3) = 0.10$ . I quote  $\lambda_{UL} = 2.3$  as the upper limit at 90% confidence."

The full description runs

For a result  $N$  the  $(1 - \alpha)$  upper limit  $\lambda_{UL}$  is chosen such that for  $\lambda = \lambda_{UL}$  or more the probability of obtaining  $N$  or less is  $\alpha$  or less.

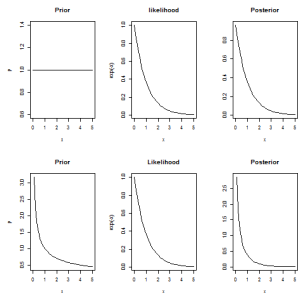
Notes on the 3 bits of small print:

- 1 The statement about  $\lambda_{UL}$  is not about a value but about the beginning of a range. 2.3 and above are ruled out.
- 2 If  $N$  is nonzero but small you may still want to quote a limit. Then you have to include the probability of getting an even smaller result.
- 3 So if a statement is true with 95% confidence, it is true with 90% confidence, etc. Also needed if the variable has discrete values.

To say " $\lambda < 2.3@90\%$  confidence" means: "The statement ' $\lambda < 2.3$ ' belongs to an ensemble of statements of which (at least) 90% are true." 

# More details

## on the Bayesian



This is easy to calculate and seems to make intuitive sense - but the 'flat prior in  $\lambda$ ' is a fudge.

Mathematically, it can't be normalised.

In reality, you don't believe it.

There is no reason to choose  $\lambda$  rather than  $\sqrt{\lambda}$ , or maybe  $\cos\lambda$ , as the variable to be flat in. These will give different results. A good analysis will try several priors (or, equivalently, several variables in which the prior is flat) and check that the result is not sensitive. ("Robustness under changes of prior")

## More details

on both

There are other useful values for zero counts. Principally  $\lambda_{UL}$  @ 95% is 3.0. For any  $N$ , Frequentist and Bayesian (with prior flat in  $\lambda$ ) quote the same values, but they mean different things. This is basically a coincidence. (It does not work for lower limits.)

# Experiments with background

Things begin to get tricky

If there is background, call it  $B$ , then  $N > 0$  does not mean an unambiguous discovery. You need  $N \gg B$  for a discovery, otherwise you quote a limit.

## 1 The Strict frequentist

Observing  $N$  gives a limit  $T_{UL}$  on the total number  $T = \lambda + B$ .

Subtract  $B$  to get the limit on  $\lambda$ .

Works fine unless  $N < B$ , or even  $N \approx B$

E.G.  $B = 2.9, N = 0$

At 95%,  $\lambda_{UL} = 0.1$       At 90%,  $\lambda_{UL} = -0.6$

The first is correct but dishonest. The second is ridiculous but correct.

## 2 The Bayesian

No problems - integrate  $\int_0^{\lambda_{UL}} e^{-\lambda+B} (\lambda + B)^N / N! d\lambda$

(But remember the ambiguity of the flat prior)

## 3 $CL_s$

## 4 Feldman Cousins

# Some technical stuff (1)

## p-values

You have some data value(s)  $x$  and a hypothesis  $H$

( $H$  is often the null hypothesis  $H_0$ .  $H_0$  is often the Standard Model)

The  $p$ -value is the probability, according to the hypothesis  $H$ , for getting a result as extreme as  $x$  (or worse).

### Notes

- 1 This 'or worse' may need care. Standard example is 1-tailed and 2-tailed Gaussian significances.
- 2 The formula is the same as for the power of a test - but  $p$  values are computed *after* seeing the data, powers are calculated *before* (In principle, anyway)
- 3 This is **not** the probability that  $H$  is true. But try telling that to journalists...



## Some technical stuff (2)

coverage

How often is a 90% CL statement true? Depends on actual value.

### Example

Poisson with true mean 4.5678

Probabilities of 0,1,2,3... counts are 0.010, 0.047, 0.108, 0.165 ...

Upper limits 2.30, 3.89, 5.32, 6.68... are false, false, true, true...

Statement has a 5.7% probability of being false, 94.3% of being true

Coverage 94.3% in this particular case.

The R code

```
dpois(0:10,4.5678)
```

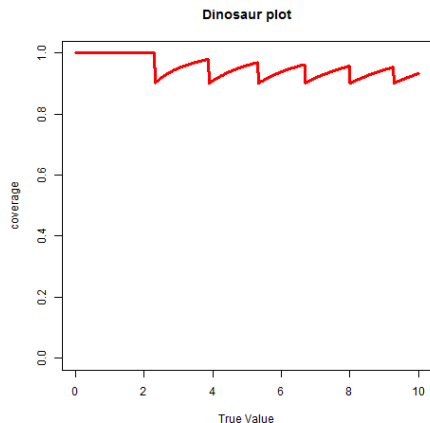
```
f<-function(x){sum(dpois(0:N,x))-0.10}
```

```
N=3
```

```
uniroot(f,c(0,10))
```

## More on coverage

Why not exactly 90%?  
Because of discrete data.  
For Frequentists, overcoverage is allowed but inefficient.  
Undercoverage is not allowed.  
For Bayesians: coverage is strictly irrelevant but gives very useful insights.



Coverage for Poisson as a function of  $\lambda$

## Some technical stuff (3)

### Realistic experiments

So far: just count events. use Poisson statistics.

In many experiments: events may be 'signal-like' and 'background-like' in varying degrees. (Especially with Neural Network or BDT outputs involved)

Simple cut-and-count loses precision.

Use Monte-Carlo to generate likelihood functions for signal and background.

In some searches (e.g. Higgs) the parameter being studied effects both the rate and the likelihood functions.

Same limit-setting principles apply

# $CL_s$

or 'modified frequentist', developed by Alex Reid and the LEP Higgs Hunters

Define  $CL_s = CL_{s+B}/CL_B$

with  $CL_{s+B} = \sum_{r=0}^N P(r; s+B)$        $CL_B = \sum_{r=0}^N P(r; B)$

Normally  $CL_B = 0.9\dots$ . Small  $CL_B$  betrays downward fluctuation.

So  $CL_s$  is bigger than  $CL_{s+B}$ , by an amount which depends sensibly on the plausibility background dipped. To reduce to 5% (or 10%...), increase  $\lambda$ .

Makes you more honest, but destroys the frequentist coverage.

## Example

Observe 3 events. Calculated background 1.2. Work with 90% limits

$p$  value  $p(\lambda) = \sum_0^3 e^{-\lambda} \lambda^r / r!$  gives 0.1 for  $\lambda = 6.68$

Straight frequentist:  $\lambda \equiv s + B$  so 90% upper limit on  $s$  is 5.48

Modified frequentist:  $p(1.2) = 0.966$ , so want  $CL_s = p(\lambda)/0.966 = 0.1$

Actually  $p(6.74) = 0.0966$ , so limit on  $s$  is 5.54.

Oops! Background recalculated. Now 5.2. Frequentist adjusts to 1.48

$p(5.2) = 0.238$ , so you want  $p(\lambda) = 0.0238$ ,  $\implies \lambda = 8.84$ , and limit 3.64

# Feldman-Cousins

Or: 'Unified Method'. A technique cunningly solves one problem by attacking another.

Confidence plot - 90% limits.

Horiz. axis. Measured  $x$ .

Vert. axis. True  $\lambda$ . Construct

$p(x; \lambda)$  for each  $\lambda$

Find  $x$  value for which

$\int_x^\infty p(x') dx' = 0.9$ . Green.

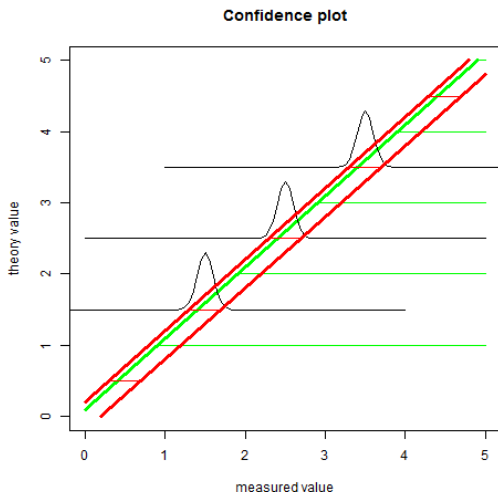
Find  $x$  values for which

$\int_{-\infty}^x p(x') dx' =$

$\int_x^\infty p(x') dx' = 0.05$ , for

central limits. Red .

Given  $x$ , get upper limit from thick green line, or central limits from thick red lines, and all is well.



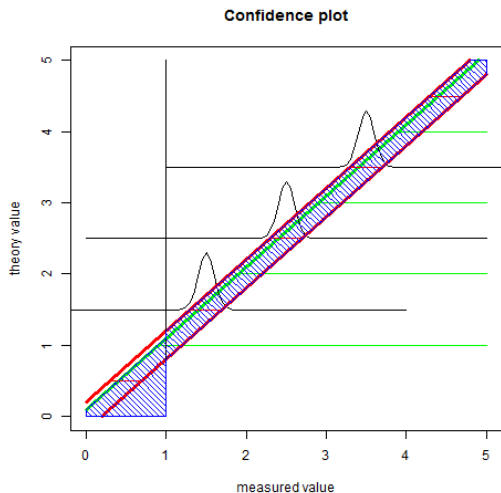
# Feldman-Cousins 2

the story continues

Real life practice

('flip-floppng'): If  $x$  is small (say, 1.0), quote an upper limit. For larger  $x$  quote a measurement.

(Means using the shaded area)  
This undercovers and is therefore evil and wrong.



# Feldman-Cousins 3

## The method

Choose any limits with

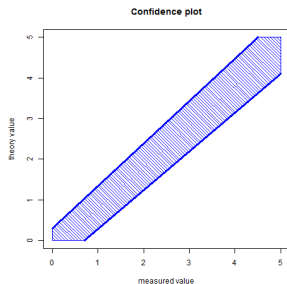
$$\int_{x_1}^{x_2} p(x; a) dx = 0.90$$

First guess: choose highest probabilities till 90% reached. (Gives shortest interval).

Minor glitch: some values - e.g.  $P(0; 3.2)$  unlikely and never get chosen - even though you would want to do so if, say, 0 events and  $B = 3.1$ .

Second guess: For a given  $a$ , rank values of  $x$  according to  $P(x; a)/P(x, a_{best})$  and choose the highest ranked till 90% reached.

For Poisson  $p(x; s + B)$ ,  $s_{best}$  is either  $x - B$ , or 0



# Feldman-Cousins 4

## How it works

For small  $x$ , get an upper limit.

For larger  $x$ , get range.

Both are good frequentist results.

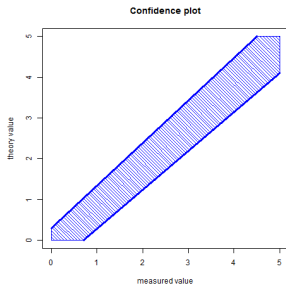
'Flip-flop' is automatic.

Need to calculate limits anew for each  $B$ .

Not a problem.

Objections raised

- 1 May quote range when you don't believe there's a real signal.  
You can live with it!
- 2 For zero events, experiments with larger backgrounds quote better limits.  
so what?





## Exercise

Your experiment detects 3 events.

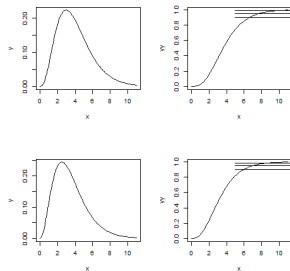
Calculate the 90%, 95% and 99% upper limits using (1) a frequentist approach (2) A Bayesian approach with a prior flat in  $\lambda$  and (3) A Bayesian approach with a prior flat in  $\sqrt{\lambda}$

# Answer...

Frequentist: define `f <- function(x){ppois(3,x)-.1}` then `uniroot(f,c(1,11))` Get 6.68, 7.75, 10.05 for 90,95,99%

Bayes: (normalised) posterior is  $e^{-\lambda}\lambda^3/6$

Can integrate algebraically - same as frequentist. Or read off graph



Bayes with prior flat in  $\sqrt{\lambda}$  - proportional to  $1/\sqrt{\lambda}$

Plot and read off - approx 6.0,.7.0,.9.1